

# Responsibility Sharing in a Public Bad Experiment\*

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## Abstract

We propose a novel mechanism to mitigate the provision of public bads in large groups. In the baseline case with centralized punishment, players choose their neighbors and having more neighbors brings benefits. Each player then decides whether to provide a public bad, which yields personal benefits but imposes costs on the entire group. A small chance exists that a player providing the public bad will be detected and punished. In the responsibility-sharing mechanism, when a player is caught providing the public bad, their neighbors are also punished. Our theoretical analysis and experimental results strongly support the effectiveness of this mechanism in promoting Pareto optimal outcomes. In addition, using another experimental treatment, we determine that the mechanism is less effective in the absence of complete feedback information.

**Keywords:** Public bads, Cooperation, Punishment, Network formation, Laboratory experiment.

**JEL codes:** C72, C92, D82, D85.

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# 1 Introduction

Global public goods such as environmental pollution, traffic congestion, overfishing, and cigarette smoking present significant challenges to societies, resulting in severe social and economic consequences such as irreversible ecological damage, worsened health outcomes, and decreased productivity. While mitigating public goods requires cooperation among individuals, such cooperation has often proven difficult based on theoretical and experimental studies (Andreoni, 1988; Isaac and Walker, 1988a,b). Nevertheless, the literature has also identified a range of effective mechanisms for addressing this issue (for overviews, see Chen, 2008; Gächter and Herrmann, 2009; Chaudhuri, 2011).<sup>1</sup>

In a world where individuals belong to various social networks, one possible approach for sustaining cooperation is allowing people to choose their interactions (Cinyabuguma et al., 2005; Page et al., 2005; Rand et al., 2011; Charness and Yang, 2014; van Leeuwen et al., 2019). Cooperation often succeeds in such settings because many players adopt *link reciprocity* strategies, forming connections with cooperative individuals and avoiding noncooperative ones (Ebel and Bornholdt, 2002; Santos et al., 2006; Hanaki et al., 2007). These studies typically assume that public goods (or goods) are local, meaning individuals are unaffected by externalities outside their communities (Cordes and Sandler, 1996). However, this assumption does not hold in circumstances concerning global public goods such as greenhouse gas emissions, in which countries cannot isolate themselves from the effects of others' actions. In such cases, *link reciprocity* may fail, as excluding noncooperative players will not prevent external harm. This raises the question of whether network structures can still enhance cooperation in global public good contexts.

This study explores mechanisms to deter the provision of global public goods in network structures. We introduce a novel responsibility-sharing mechanism, in which individuals form networks and only share responsibility with direct connections. The mechanism assumes that agents are motivated to connect with others, driven by intrinsic factors, such as regional free-trade agreements, or external incentives, such as government promotion opportunities based on the number of connections.<sup>2</sup>

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<sup>1</sup>Studies on the public goods problem can be integrated into the framework of public goods game (Shitovitz and Spiegel, 2003).

<sup>2</sup>Exogenous sources can include information or technology sharing among interconnected individuals or companies (Bala and Goyal, 2000; Jackson and Watts, 2002; Harrison and Muñoz, 2008; Goeree et al., 2009; Falk and Kosfeld, 2012; Jackson and Zenou, 2015).

We introduce a two-stage game. In the first stage, each player proposes links to others in the group, with a neighbor only forming if both parties propose it and having more neighbors brings benefits. In the second stage, players observe the network structure and decide whether to provide a public bad, which offers personal benefits but imposes costs on society. Crucially, the benefits of the public bad are less than the total costs it generates, creating a social dilemma. While abstaining from providing the public bad is socially optimal, providing it remains the dominant strategy for all players. We assume complete information, meaning players learn the choices made by others at the end of the game.

We introduce centralized punishment (CP) and responsibility-sharing (RS) mechanisms into the game. The CP mechanism reflects real-world systems, in which each player faces a small, exogenous probability of detection and subsequent punishment if caught providing the public bad. The RS mechanism extends the CP model by not only imposing punishment on the detected player, but also on their direct neighbors within the network. Our focus is investigating behavior and cooperation efficiency within the RS mechanism, with the CP mechanism serving as a benchmark to control for factors unrelated to RS.

Theoretically, both the CP and RS mechanisms can support full-cooperation and noncooperation outcomes as subgame-perfect equilibria in the infinitely repeated game. In the noncooperation equilibrium, all players provide the public bad. Under CP, players form a large network where everyone is connected, while the network is smaller under RS, with each player having one neighbor at most. In both the mechanisms, the full-cooperation equilibrium is efficient and yields higher payoffs for all players compared with the noncooperation equilibrium. This outcome is achieved when players form all links and refrain from providing the public bad. However, based on the basin of attraction theory (Dal Bó and Fréchette, 2011), we find that the full-cooperation equilibrium is more attractive in RS than in CP when the discount factor is sufficiently large. In the RS mechanism, the noncooperation equilibrium becomes less appealing due to the loss of network benefits, as a larger network cannot be supported in this equilibrium.

In our laboratory experiment, we use a between-subjects design to compare the behavior of the two mechanisms in CP and the RS treatments. In both treatments, participants are randomly assigned to eight-player groups and play the respective game

in an infinitely repeated format, with random termination determined by a continuation probability of 0.97 after each period. The experimental results reveal a stark contrast between the two treatments. Nearly all groups in the RS treatment converge to full cooperation with complete networks, while no groups in the CP treatment achieve this outcome. Specifically, in the RS treatment, all participants, especially cooperative ones, are more likely to strategically form links with other cooperators than with noncooperators, drawing on reputations established from past behavior. This sorting behavior encourages noncooperative participants to switch to cooperation when they realize it is more profitable, resulting in full cooperation within the group. In contrast, participants' link decisions in the CP treatment do not depend on others' past behavior, resulting in a majority continuing to provide the public bad.

Moreover, assuming complete information among all players may be unrealistic. To address this, we introduce a private information (PI) treatment, which closely mirrors the RS treatment but with the key difference that players' public bad actions remain private. Players also have the option of revealing their actions at a cost. In the experiment, similar to the RS treatment, half of the PI groups achieve full cooperation, while the other half exhibit a pattern more like the CP treatment. Analyses reveal that successful cooperation in PI groups is driven by a higher disclosure rate, which establishes an environment that is more closely aligned with the RS treatment. Notably, participants rarely disclose their actions once full cooperation is achieved, suggesting that while information disclosure is crucial for fostering initial cooperation, it becomes less essential once cooperation is established.

Finally, while this is the first study to apply responsibility sharing to improve cooperation in a global public bad game, the concept has been widely employed in practical management, particularly in local game settings across fields such as finance, accounting, environmental protection, and import-export activities.<sup>3</sup> By extending these principles to the global context, the RS mechanism offers practical solutions to mitigate

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<sup>3</sup>For instance, microfinance institutions in developing countries employ group-lending contracts, where loans are provided collectively to a group of low-income borrowers, with the entire group held liable if any member defaults on repayment (Morduch, 1999; Attanasio et al., 2015). Similarly, when banks identify a higher proportion of bad borrowers in a pool, they tighten lending standards, making it more difficult for all entrepreneurs in the pool to secure loans (Fishman et al., 2024; Fan, 2021). For more examples, see Oei (2017), Oei and Ring (2018) and The Pie News (2023).

global public goods, which are easily implementable when a centralized institution can detect and punish misbehavior, and when individuals have sufficient externally provided or naturally occurring incentives to form networks. For example, to combat corruption among government officials, authorities could implement policies that reward individuals with more connections and a willingness to share responsibility for corruption within their networks. Similarly, in international climate agreements, countries could receive enhanced trade benefits and developmental opportunities by forming more connections, contingent on substantial cooperation in sharing responsibility for potential violations.

## 2 Related Literature

This study is primarily related to three streams of literature. First, it contributes to enhancing the understanding of cooperative behavior within network structures. Previous research has predominantly focused on local public goods and goods, where individuals are unaffected by the impacts of these goods or goods beyond their immediate communities. For example, Cinyabuguma et al. (2005) find that in the voluntary contribution mechanism, groups exhibit nearly complete cooperation when players face potential exclusion based on majority voting. Similar results under different regrouping rules are observed by Page et al. (2005) and Charness and Yang (2014). In network structures, where players can autonomously choose their interaction partners, Rand et al. (2011) observe increased cooperation when players have the ability to frequently reconfigure links with others.<sup>4</sup> The success of cooperation in these studies is often attributed to the use of *link reciprocity*, which means players strategically form links with cooperative players and sever ties with noncooperative players (Ebel and Bornholdt, 2002; Santos et al., 2006; Hanaki et al., 2007; Shirado et al., 2013; Gallo and Yan, 2015; Riedl et al., 2021).<sup>5</sup>

In a closely related study, Riedl et al. (2016) investigate a two-stage game similar to

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<sup>4</sup>He and Zou (2024) explore public goods provision within a network formation game, finding that cooperative behavior is highly sensitive to the costs of establishing links.

<sup>5</sup>The use of *link reciprocity* can be influenced by several factors. For example, van Leeuwen et al. (2019) find that players with central positions in a network are able to maintain connections with others even when their cooperation levels are relatively low. For more related findings, see Gächter and Thöni (2005), Ahn et al. (2009), Fowler and Christakis (2010), Brekke et al. (2011), Gracia-Lázaro et al. (2012), and Li et al. (2018).

ours, with the distinction that their second stage is a weakest-link coordination game. In their design, each player selects a single effort level that applies to all neighbors and payoffs are determined by the lowest effort level within their endogenously formed network. They find that players widely adopt *link reciprocity* to facilitate coordination in large groups with 24 players. Our study also relates to Riedl et al. (2016) in several noteworthy ways. First, we show that in endogenously formed networks, participants do not reach socially optimal outcomes in our global public bad game, in contrast to the findings of Riedl et al. (2016) for weakest-link games. A potential explanation is that attaining socially optimal outcomes in a public bad game requires players to choose dominated actions, whereas this is not the case in weakest-link games. Therefore, the public bad game environment is arguably more demanding. However, we identify a very effective RS mechanism through which socially optimal outcomes can be sustained in our setting. Finally, it is worth noting that the combination of responsibility sharing with probabilistic centralized punishment in a global public bad game establishes an environment that resembles a weakest-link setting, insofar as an individual's payoff may depend on the behavior of the least cooperative neighbors in the network. Nevertheless, the two settings are conceptually distinct: in our RS setting, the socially optimal action is strictly dominated in the stage game and can only be sustained through repeated-game incentives, whereas in the weakest-link game of Riedl et al. (2016) the socially optimal action profile constitutes a stage-game Nash equilibrium.

Second, our study contributes to the literature on the role of punishment in fostering collective cooperation. Previous studies have shown that under decentralized punishment institutions —where individuals can punish others— cooperators frequently punish free-riders, even when doing so entails personal costs (Fehr and Gächter, 2000; Masclet et al., 2003; Denant-Boemont et al., 2007; Nikiforakis and Normann, 2008; Nikiforakis, 2008; Xiao and Houser, 2011; Cason and Gangadharan, 2015; Boosey and Isaac, 2016).<sup>6</sup> While

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<sup>6</sup>Several studies explore the relationship between network structures, that determine who can punish whom and decentralized punishment (Carpenter et al., 2012; Leibbrandt et al., 2015; DeAngelo and Gee, 2020). The general finding is that complete networks tend to yield the most efficient outcomes. Moreover, Leibbrandt et al. (2015) find that network configurations has a more significant influence than punishment capacities on determining the level of public good provision. DeAngelo and Gee (2020) further demonstrate that when networks are endogenously formed, peer monitoring proves more effective than group monitoring.

such behavior compels potential free-riders to cooperate in order to avoid punishment, the misuse of punishment rights may lead to antisocial punishment and reduce overall cooperation efficiency (Herrmann et al., 2008; Nikiforakis and Normann, 2008; Chaudhuri, 2011). Next, under CP institutions, where an authority imposes penalties, cooperation tends to improve more substantially when these institutions are endogenously chosen rather than exogenously imposed (Tyran and Feld, 2006; Bó et al., 2010; Puttermann et al., 2011). This effect is particularly pronounced when exogenously imposed punishment institution lack sufficient deterrence, such as in case of low detection rates or weak punishment severity (Anderson and Stafford, 2003; Tyran and Feld, 2006; Bó et al., 2010; Andreoni and Gee, 2012; Kamijo et al., 2014). The RS mechanism proposed in this study enhances the deterrent power of CP without introducing additional management costs, provided that individuals are sufficiently motivated to form connections with others.

Finally, our study extends the concept of community enforcement, which has long been recognized as a central mechanism for sustaining social cooperation (Ostrom, 1990; Ostrom et al., 1992; Greif, 1989, 1993; Greif et al., 1994; Dixit, 2003). Previous literature primarily conceptualizes community enforcement as deterring noncooperative behavior through the threat of widespread cooperation breakdown across society (Kandori, 1992; Ellison, 1994; Takahashi, 2010; Wolitzky, 2013; Ali and Miller, 2014; Deb, 2020). In this framework, the victim of a defector typically adopts a grim trigger strategy wherein lacking a direct means of punishing the defector, they withdraw cooperation from all members of the community as an indirect form of punishment. This response spreads noncooperation throughout the community and ultimately punishes the original defector. In contrast, the community enforcement induced by the RS mechanism in our study differs in two notable ways. First, punishment is confined to the defector's immediate network, ensuring that uninvolved parties outside this network are not subject to collateral punishment. Second, unlike traditional community enforcement, where the victim assumes the role of enforcer, the RS mechanism assigns this role to a centralized institution. This distinction enables more targeted punishment and reduces the risk of escalation into widespread noncooperation.<sup>7</sup> Finally, although the victim could, in principle, punish the entire

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<sup>7</sup>It can therefore be interpreted as a form of community enforcement through specialized enforcement mechanisms. Several studies have compared the relative roles of these two forms of enforcement in sustaining cooperation (Masten and Prüfer, 2014; Aldashev and Zanarone, 2017; Acemoglu and Wolitzky, 2020).

community by withholding cooperation, our results demonstrate that the availability of targeted punishment under the RS mechanism renders such strategies rarely used.

## 3 Theory

In this section, we present a model with costless links, which is consistent with the parameters used in the experiment. In Online Appendix A, we extend the model to a generalized version that allows for positive link costs.

### 3.1 Model

We construct a public bad network game. In the first stage, players endogenously form a network, where having more neighbors increases individual payoffs. In the second stage, players simultaneously decide whether to provide a public bad that imposes costs on all other players. In the baseline model, all players face a positive possibility of detection and those detected for choosing the public bad are punished. In the main model, we introduce a RS mechanism under which the neighbors of a detected provider are also punished.

#### 3.1.1 Public bad network game

Eight players participate in the game, indexed as  $N = \{1, 2, \dots, 8\}$ . In the first stage, players simultaneously and independently choose who to propose links to. Player  $i \in N$  chooses a vector of proposals  $\mathbf{g}_i = (g_{i1}, g_{i2}, \dots, g_{i8})$ , where  $g_{ij} = 1$  if  $i$  proposes a link to player  $j$  and  $g_{ij} = 0$  otherwise. Let  $\mathbf{g} = (\mathbf{g}_1, \mathbf{g}_2, \dots, \mathbf{g}_8)$  denote the network of all players' proposals, which forms a directed graph. Next, we adopt the two-sided link formation framework proposed by Jackson and Wolinsky (1996), where mutual consent is required to establish a neighbor relationship. In other words, players  $i$  and  $j$  become *neighbors* if and only if they propose a link to one another. Denote  $N_i(\mathbf{g}) = \{j \in N | g_{ij}g_{ji} = 1, j \neq i\}$  as the set of player  $i$ 's neighbors, and let  $|N_i(\mathbf{g})|$  denote the number of  $i$ 's neighbors. Suppose the cost of each link proposal is  $c$ . For the purposes of our experiment, we assume  $c = 0$ , which simplifies participants' decision-making and aligns with a common assumption in the experimental literature on network games (Rand et al., 2011; Riedl

et al., 2016; Teteryatnikova and Tremewan, 2020; Goyal et al., 2021).<sup>8</sup> In addition, we introduce a scale effect within the network game, wherein each neighbor yields a benefit of 20. In other words, player  $i$  receives an additional benefit of  $20|N_i(\mathbf{g})|$  by establishing  $|N_i(\mathbf{g})|$  neighbors.

In the second stage, after observing the network  $\mathbf{g}$  formed in the first stage, each player simultaneously chooses an action,  $a_i \in \{x, y\}$ . Let  $\mathbf{a} = (a_1, a_2, \dots, a_8)$  denote an action profile of the second stage of the game. Choosing  $x$  represents not providing the public bad, while choosing  $y$  represents providing it. A player who chooses  $x$  obtains a benefit of 50 and imposes no externalities on others, while a player who chooses  $y$  earns 100 but imposes a cost of 15 on every other group member. Therefore, although  $y$  maximizes individual payoffs,  $x$  is socially efficient.

Player  $i$ 's final payoff is determined by two components, the gain and the loss (denoted as  $U_{i,gain}$  and  $U_{i,loss}$ , respectively). The gain is calculated as the sum of the benefits derived from the network and the action in the second stage, which is represented as follows:<sup>9</sup>

$$U_{i,gain} = 20|N_i(\mathbf{g})| + 50(100), \text{ if } a_i = x(y). \quad (1)$$

The loss equals the total costs imposed by the other players who choose  $y$  in the group, which is represented as follows:

$$U_{i,loss} = 15 \sum_{k \in N, k \neq i} \mathbb{I}\{a_k = y\}. \quad (2)$$

Finally, player  $i$ 's payoff is the difference between the two components, as follows:

$$U_i(\mathbf{g}, \mathbf{a}) = U_{i,gain} - U_{i,loss}. \quad (3)$$

A key assumption of the public bad network game is that each additional neighbor yields a fixed benefit of 20. This reflects network economies of scale, whereby players derive direct benefits from expanding their connections. For example, a larger pool of Uber drivers increases consumer adoption of the platform, subsequently improving individual drivers' earnings. Similarly, firms located in larger industrial parks benefit from reduced transportation costs by leveraging synergies within the industrial chain.<sup>10</sup>

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<sup>8</sup>In the generalized model presented in Online Appendix A, we relax the assumption of  $c = 0$  in two ways. (1) We introduce a positive cost for forming a neighbor through bilateral proposals, and (2) in Online Appendix A.3, we examine the case of costly unilateral link proposals ( $c > 0$ ).

<sup>9</sup>We assume linearity in combining the benefits arising from the scale effect of the network and the payoffs associated with actions  $x$  or  $y$ .

<sup>10</sup>Various studies employ different formulations of network economies of scale. For example, in some

### 3.1.2 Baseline model: CP with random detection

The baseline model incorporates CP with random detection into the public bad network game. Specifically, following the second-stage action choices, a central institution randomly detects each player's actions with a fixed detection probability. If a player who chose action  $y$  is detected, they incur a fixed punishment, setting their gain component of the payoff to 0, while their loss component remains unaffected.<sup>11</sup> Conversely, if a player who chose  $y$  is not detected or if a player who chose  $x$  is detected, no punishment is imposed, and the gain and loss components of their payoffs remain unaffected. In summary, the gain component of player  $i$ 's payoff in the baseline model can be described as follows:

$$U_{i,gain} = \begin{cases} 20|N_i(\mathbf{g})| + 50(100), & \text{if } i \text{ is not detected to choose } y \text{ and } a_i = x(y) \\ 0, & \text{if } i \text{ is detected to choose } y \end{cases}. \quad (4)$$

We assume that the detection process is exogenous and independent across players, with a fixed probability of 15%. This relatively low detection rate implies that choosing action  $x$  is not a dominant strategy.

### 3.1.3 Main model: RS mechanism

We extend the baseline model by introducing the RS mechanism into the main model. Under this mechanism, if player  $i$  is detected choosing action  $y$ , then player  $i$  is punished with a zero gain and all of their neighbors also incur the same punishment. In other words, the consequences of providing the public bad extend to player  $i$ 's immediate network. Similarly, the RS mechanism's punishment only affects the gain component of the payoff, settings a player's payoff increases proportionally with neighborhood size (Riedl et al., 2016). Notably, we demonstrate that the equilibrium properties derived below remain robust, employing a setting akin to that of Riedl et al. (2016).

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<sup>11</sup>An alternative approach to model the punishment payoff is to incorporate a fixed deduction that players receive when they are punished. For the sake of simplicity, we choose to assign a fixed status quo payoff of zero to players upon being punished.

which is presented as follows:

$$U_{i,gain} = \begin{cases} 20|N_i(\mathbf{g})| + 50(100), & \text{if neither } i \text{ nor any player in } N_i(\mathbf{g}) \text{ is detected} \\ & \text{to choose } y, \text{ and } a_i = x(y) \\ 0, & \text{if } i \text{ or at least one player in } N_i(\mathbf{g}) \text{ is detected} \\ & \text{to choose } y \end{cases} . \quad (5)$$

## 3.2 Equilibrium analysis

This section begins by examining the equilibrium of the one-shot game for baseline and main models. Owing to the dynamic nature of the two-stage game, we adopt the subgame-perfect equilibrium as the solution concept.<sup>12</sup> We eliminate weakly dominated strategies as a refinement of multiple equilibria in both models. We then analyze the equilibrium of the infinitely repeated game. In Online Appendix A, we demonstrate that all the equilibrium properties presented in this section hold in the generalized model.<sup>13</sup>

### 3.2.1 Equilibria of the one-shot game

We next derive the equilibrium for the one-shot game of the baseline model described in Section 3.1.2, and the main model described in Section 3.1.3.<sup>14</sup> The one-shot game includes two stages, which we analyze using backward induction. Proposition 1

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<sup>12</sup>Although the network formation process in our first-stage game follows Jackson and Wolinsky (1996), we do not adopt their solution concept of pairwise stable networks because this study focuses on an infinitely repeated game in which the primary interests lie in the dynamic strategies underlying the formation of neighbor relationships.

<sup>13</sup>The generalized model introduces positive link cost for each neighbor formed, and all equilibrium properties continue to hold when this cost falls below a certain threshold. A cost higher than this threshold reduces the equilibrium set of the one-shot game, producing equilibrium outcomes with fewer neighbors in the first stage, without affecting the dominant choice of  $y$  in the second stage. In the alternative setting of costly unilateral proposal discussed in Online Appendix A.3, we demonstrate that the equilibrium properties of the one-shot game are similar to those in Propositions 2 and 3, except that unilateral proposals no longer exist in equilibrium. Moreover, while the “elimination of weakly dominated strategies” used in Section 3.2.2 is not useful in the one-shot game, the equilibria of the infinitely repeated game remain valid.

<sup>14</sup>We do not derive the equilibrium of the public bad network game without CP introduced in Section 3.1.1, as it is straightforward that proposing to all other players and providing the public bad are dominant strategies.

characterizes the equilibrium strategies in the second-stage subgame.

**Proposition 1.** *For baseline and main models, given any network  $\mathbf{g}$ , the unique equilibrium strategy for the second-stage subgame is  $a_i = y$  for all  $i \in N$ .*

Proposition 1 shows that the unique equilibrium strategy of the second-stage subgame for any network formed in the first stage in baseline and main models is all players choosing action  $y$ . As shown in Online Appendix B, this is because the marginal benefit of choosing  $y$  always exceeds the expected punishment given the low detection rate.

Next, we derive the equilibrium for the one-shot game. Proposition 2 presents the equilibrium properties of the baseline model.

**Proposition 2.** *For every subgame-perfect equilibrium of the baseline model,  $(\mathbf{g}, \mathbf{a})$ , the following properties hold for all  $i, j \in N$  with  $i \neq j$ : (i)  $g_{ij} = g_{ji}$  and (ii)  $a_i = y$ .*

The first condition of Proposition 2 imposes no constraints on the equilibrium network structure, with the exception of the requirement that all link proposals in the network should be bilateral. The rationale behind this condition is that if a player unilaterally proposes a link to another player, it is always profitable for the latter to form a bilateral link by proposing back, irrespective of the choices made by both players in the second stage. Examples of equilibrium networks include the *complete* network, wherein each player proposes links to all others, and the *empty* network, where no links exist.

Proposition 3 presents the equilibrium properties of the main model.

**Proposition 3.** *For every subgame-perfect equilibrium of the main model,  $(\mathbf{g}, \mathbf{a})$ , the following properties hold for all  $i, j \in N$  with  $i \neq j$ : (i)  $|N_i(\mathbf{g})| \leq 1$ , (ii) if  $|N_i(\mathbf{g})| < 1$ ,  $g_{ji} = 0$ , and (iii)  $a_i = y$ .*

The first condition of Proposition 3 states that, in equilibrium, each player has at most one neighbor. This restriction arises because having more neighbors increases the risk of punishment in the main model and all players choose action  $y$  in the second stage. Moreover, the expected payoff initially rises with an increased number of neighbors but eventually declines. Consequently, we find that the optimal number of neighbors to maximize the expected payoff is one. The second condition asserts that, in equilibrium, a player should not unilaterally propose a link to someone with fewer than one neighbor. Otherwise, the latter player has an incentive to propose back to receive a higher payoff.

One example of an equilibrium network is the *1-regular* network in which all players have exactly one neighbor. Specifically, we divide the eight players into four disconnected pairs and players within each pair are neighbors to one another. Moreover, the empty network can also be an equilibrium network in the main model.

### 3.2.2 Eliminating weakly dominated strategies

In the baseline and main models, the unique equilibrium in the second-stage subgame is that all players provide the public bad. However, various network structures can be sustained in the first-stage equilibrium. In the baseline model, players can form any network structure without unilateral proposals, while in the main model, players can have one neighbor at most.

Given the multiplicity of equilibria in the one-shot game in both models, we introduce a refinement that eliminates weakly dominated strategies. Imposing this refinement enables us to focus on the equilibria in which players propose more links to one another under zero link costs. To illustrate, consider the equilibrium strategy of proposing zero links and choosing action  $y$  in the baseline model. This strategy is weakly dominated by the strategy of proposing links to all other players and choosing  $y$  because the latter yields a strictly higher payoff for the player when at least one other player proposes a link back, yielding a benefit from establishing more neighbors. The payoff remains unchanged when other players do not propose links to the player, given the zero link cost. Proposition 4 illustrates that, by imposing this refinement in the baseline model, the only surviving equilibrium is characterized by players forming a complete network and choosing action  $y$ . Conversely, the equilibrium in the main model involves players forming a 1-regular network and choosing action  $y$ .

**Proposition 4.** *After eliminating weakly dominated strategies, the baseline model has a unique, undominated subgame-perfect equilibrium,  $(\mathbf{g}, \mathbf{a})$ , where for all  $i, j \in N$  with  $i \neq j$ : (i)  $g_{ij} = 1$  and (ii)  $a_i = y$ , and the undominated subgame-perfect equilibrium of the main model,  $(\mathbf{g}, \mathbf{a})$ , satisfies the following properties for all  $i \in N$ : (i)  $|N_i(\mathbf{g})| = 1$  and (ii)  $a_i = y$ .*

### 3.2.3 Equilibria of the infinitely repeated game

In this section, we examine the infinitely repeated game of baseline model and main models with a discount factor  $\delta$  and derive the corresponding equilibrium. First, we demonstrate that adopting the undominated equilibrium strategy of the one-shot game in each period constitutes a subgame-perfect equilibrium in the infinitely repeated game. We refer to these equilibria as *noncooperation*, as all players choose action  $y$  in these equilibria.

**Proposition 5** (Noncooperation equilibrium). *The strategy in which, for every period and history, each player consistently plays the equilibrium strategy of the one-shot game in the baseline and the main model outlined in Proposition 4, constitutes a subgame-perfect equilibrium of the infinitely repeated game of each respective model.*

In addition to the noncooperation equilibria, we further confirm that a full-cooperation outcome in which all players form a complete network and choose action  $x$  in each period can be supported by a subgame-perfect equilibrium of the infinitely repeated game with the following grim trigger strategy.<sup>15</sup>

**Definition 1.** *The grim trigger strategy is defined as follows:*

- (i) *For baseline and main models, in each period from the first period, each player proposes links to all other players in the first stage and chooses  $x$  in the second stage;*
- (ii) *In the baseline model, if any player deviates from (i) in period  $t \geq 1$ , all players propose links to all other players in the first stage and choose  $y$  in the second stage in every period after  $t$ ;*
- (iii) *In the main model, if any player deviates from (i) in period  $t \geq 1$ , all players propose a single link, forming a fixed 1-regular network in the first stage and choose  $y$  in the second stage in every period after  $t$ .*

**Proposition 6** (Full-cooperation equilibrium). *When  $\delta \geq 0.13$ , the grim trigger strategy constitutes a subgame-perfect equilibrium of the infinitely repeated game of the baseline*

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<sup>15</sup>Other outcomes in which players choose action  $x$  but establish different networks may also be supported in the infinitely repeated game. This study focuses on a complete network as it is the most focal network due to its efficiency.

*model. When  $\delta \geq 0.06$ , the grim trigger strategy constitutes a subgame-perfect equilibrium of the infinitely repeated game of the main model.*

Proposition 6 indicates that full cooperation can be sustained in the infinitely repeated game if players revert to an equilibrium of the one-shot game that survives the elimination of weakly dominated strategies following any deviation, as outlined in Proposition 4. Specifically, players in the baseline model are threatened by the risk of loss resulting from all other players choosing action  $y$ , whereas players in the main model not only incur losses when other players choose  $y$ , but also face the risk of losing neighbors, as they form a 1-regular network and choose  $y$  as a consequence of deviation. Notably, the full-cooperation equilibrium exists for both models. The detailed proof is presented in Online Appendix B.

### 3.2.4 Comparing the likelihood of cooperation

Although we have demonstrated that both noncooperation and full cooperation can be equilibrium in the infinitely repeated games of baseline and main models, the question of which equilibrium is more likely to occur remains unanswered. To address this, we employ the basin of attraction criterion proposed by Dal Bó and Fréchette (2011). The authors argue that the decision of whether to adopt a cooperative strategy in an infinitely repeated cooperation game depends on various factors such as game parameters and beliefs regarding the probability of other opponents adopting the cooperative strategy. Their findings reveal a negative correlation between the probability of evolving toward cooperation and the size of the basin of attraction in repeated-game strategies.

By considering these insights, we compare the likelihood of full cooperation between the two models based on their basin of attraction of the noncooperation strategy. The expected payoff of the full-cooperation equilibrium is 190 for each player in every period, regardless of the model type. However, the noncooperation equilibrium yields a payoff of 134 in the baseline model, whereas it only yields a payoff of 16.7 in the main model. This discrepancy arises from the fact that players in the noncooperation equilibrium of the main model can, at best, establish links with only one other player in the group. In contrast, in the baseline model, players benefit from establishing links with more players without being affected by their choices regarding the public bad. Intuitively, the lower payoff associated with the noncooperation equilibrium renders the full-cooperation equilibrium

relatively more appealing in the main model compared with the baseline model.

Formally, we focus on two equilibrium strategies of the infinitely repeated game, encompassing the grim trigger strategy employed in the full-cooperation equilibrium, and the noncooperation strategy where players repeatedly play a one-shot game equilibrium that survives the elimination of weakly dominated strategies, as outlined in Proposition 5.<sup>16</sup> We calculate the size of the noncooperation strategy's basin of attraction respectively for baseline and main models, which is the probability that a player assigns to each opponent using the grim trigger strategy such that the player is indifferent between playing the two strategies. A larger basin of attraction indicates that a player must have a stronger belief in the opponents' cooperation to opt for the cooperation strategy (grim trigger), indicating a lower likelihood for cooperation to emerge. As a result, we arrive at the following proposition.<sup>17</sup>

**Proposition 7** (Basin of attraction). *If  $\delta \geq 0.55$ , the size of the noncooperation strategy's basin of attraction in the main model is smaller than that of the baseline model.*

This proposition indicates that in our model, where the discount factor is assumed to be sufficiently high, full cooperation is more likely to arise in the main model compared with the baseline model.

## 4 Experimental Design, Procedures, and Hypotheses

### 4.1 Experimental design

In the experiment, we employ a between-subjects design to implement three treatments: the Centralized Punishment treatment (CP), the Responsibility-Sharing treatment (RS), and the Private Information treatment (PI). In the CP and RS treatments, the game closely follows the baseline and main models, respectively. Each participant's link proposals and actions in CP and RS treatments are public information that can be observed by all other participants in the group. Additionally, in the PI treatment,

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<sup>16</sup>Proposition 5 allows players to establish different neighbors across periods and to propose unilateral links in the noncooperation equilibrium. However, for calculation simplicity, we restrict attention to noncooperation strategies in which players form a fixed 1-regular network over periods without unilateral proposals.

<sup>17</sup>See Online Appendix B for the detailed basin of attraction calculation and proof of Proposition 7.

participants play the game based on the main model, but their actions in the second stage are private and undisclosed, except when they are detected for choosing the public bad action. However, participants can voluntarily disclose their actions to the group at an additional small cost.

The CP and RS treatments are designed to identify the effect of the RS mechanism. The PI treatment is introduced for two main reasons. First, it enables us to decompose the factors that may contribute to improving cooperation in the RS treatment. Second, it mirrors real-world settings more closely and therefore has greater empirical relevance. In many contexts, whether an agent has engaged in harmful behavior (a public bad) is private information to others in society. Examples include illicit tax evasion, pollutant emissions, or food safety violations, which are often conducted discreetly, sometimes without the knowledge of close partners or collaborators. However, in many cases, agents can voluntarily disclose their actions to the public, albeit at a cost. For example, chemical companies may publish emissions data on their websites to demonstrate environmental responsibility and food manufacturers can display images of their production facilities as evidence of compliance with safety standards.

In all three treatments, we conduct the infinitely repeated game in the laboratory through a process known as random termination (Roth and Murnighan, 1978). At the end of each period, there is a fixed and known probability of 0.97 that the game will continue for another period. This implies that the expected number of periods is approximately 33.<sup>18</sup> Fréchette and Yuksel (2017) demonstrate that the game under random termination, characterized by a continuation rate of  $r$ , is theoretically equivalent to an infinitely repeated game for a risk-neutral player, with a discount factor of  $r$ .

At the beginning of the experiment, participants are randomly assigned to a group of eight players and play the game with the same partners throughout the experiment. Each participant is assigned a fixed alphabetical player ID (A–H).<sup>19</sup>

#### 4.1.1 CP and RS treatments

In the CP and RS treatments, participants play the public bad network game in each round. In the first stage of each round, they simultaneously and independently decide

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<sup>18</sup>The expected number of periods is calculated by  $\frac{1}{1-0.97}$ .

<sup>19</sup>As shown in the experimental screenshots in Online Appendix C, each player is represented by a node in the network graph, with player IDs displayed nearby.

who to propose links to. In the second stage, after observing the formed network (including bilateral and unilateral link proposals), participants simultaneously choose between actions  $x$  and  $y$ . Once all decisions are made, each participant's second-stage action is independently detected with a probability of 15%. In CP, only participants detected to have chosen  $y$  are punished, while in RS, the detected participants and their neighbors are punished. Punishment reduces the gain component of the affected player's payoff to zero.

In both treatments, participants begin the first stage of round one without any prior history, only an empty network displaying player IDs. At the end of each round, they receive feedback on that round, including the formed network (with established neighbors and unilateral link proposals), the actions and payoffs of each player in the group, and the detected players. From round two onward, participants can review all feedback from previous rounds.<sup>20</sup> Since player IDs remain fixed throughout the experiment, participants can track others' past behavior, enabling reputation formation.

#### 4.1.2 PI treatment

In the PI treatment, participants repeatedly play the game of the main model with random termination, similar to the RS treatment. However, in contrast to the RS treatment, participants' choices between actions  $x$  and  $y$  are private information in the PI treatment, and they have the option to choose whether to disclose their actions to all group members.

In each round, following the two decision stages, the computer randomly detects each participant's action with a fixed probability of 15%. After this, participants first receive feedback on that round, which includes the formed network, the participants who are detected for choosing  $y$ , the total number of group members who have chosen the public bad action, and their own payoffs.<sup>21</sup>

Following the above feedback, an additional action-disclosure stage is implemented at the end of each round. In this stage, participants decide whether to disclose their

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<sup>20</sup>Figure A.1 in Online Appendix C provides a screenshot of the decision-making page of the first stage. On the left of this screen, the feedback from the most recent round is displayed as a default and participants can access the feedback from any other previous round.

<sup>21</sup>The disclosure of detected  $y$  mirrors real-life practices, where public bad actions are typically made public once detected. For example, regulatory agencies publish reports on restaurants that fail to meet food safety standards, and tax authorities disclose identified cases of tax evasion.

actions from the current round at a cost of five units. If disclosed, their actions and payoffs become publicly observable to all group members at the start of the next round; otherwise, they remain hidden and are shown as blank. As in CP and RS treatments, participants in the PI treatment have access to all available information from previous rounds at the beginning of each round. The key difference is that in the PI treatment, the history excludes the actions and payoffs of participants who chose not to disclose them.

## 4.2 Procedures

Our experiment was conducted at the Shanghai University of Finance and Economics in 2022. Participants were recruited from the subject pool of the Economic Lab using Ancademy.<sup>22</sup> Each participant was enrolled in only one of the three treatments. We conducted six sessions with four groups in each session. To control for session effects, two treatments were conducted simultaneously in each session. We conducted eight groups of eight participants for each of the three treatments. A total of 192 students participated in the experiment, most of whom were undergraduate students with various majors.

The participants interacted via computer terminals and we programmed the experiment using z-Tree (Fischbacher, 2007). Upon arrival, participants were randomly seated in the laboratory. At the beginning of the experiment, they were required to read the instructions displayed on the computer screen and answer all control questions correctly. At the end of the experiment, participants answered a questionnaire of demographic information. The experimental instructions and screenshots are provided in Online Appendix C.

To simplify the between-treatment comparisons, we performed random terminations before running the sessions and chose the same set of period numbers across the three treatments. As a result, the randomly realized period number for the eight independent groups for all treatments are 17, 27, 30, 32, 39, 40, 43, and 47.

The participants earned currency points in the experiment, which were calculated as the sum of all points accumulated across all rounds, and the exchange rate is 80 points = 1 Chinese *yuan* (CNY). On average, participants earned 78 CNY (equivalent to about 11 USD), including a show-up fee of 15 CNY (about 2 USD). Each session lasted between 75 and 100 minutes.

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<sup>22</sup>Ancademy is a platform for social sciences experiments.

### 4.3 Hypotheses

Our primary goal is to examine the effects of the RS mechanism in our public bad network game, for which we propose a hypothesis regarding the behavioral differences in CP and RS treatments. Then, we propose a hypothesis for the effect of the information environment regarding the behavioral differences in RS and PI treatments.

According to the theoretical analysis in Section 3.2, although noncooperation and full cooperation can be sustained as equilibrium in CP and RS, the basin of attraction theory indicates that full cooperation is more likely to arise in RS than in CP. Therefore, we propose the following hypothesis:

**Hypothesis 1.** *The rate of  $y$  choices is lower in the RS treatment than in the CP treatment.*

Next, we develop a hypothesis regarding the effect of the information environment. First, as disclosing information in PI incurs cost, we conjecture that players intending to choose action  $y$  (referred to as “defectors”) will not reveal their decisions to avoid damaging their reputations. In contrast, players intending to choose action  $x$  (referred to as “cooperators”) have an incentive to disclose their decisions because it can establish a favorable reputation (Kirchsteiger et al., 2024). Assuming that cooperators consistently choose to disclose their actions, we anticipate that the pattern of public bad action choices in PI will resemble that of RS, as a player’s decision not to disclose their choice unequivocally signals a defector type. In summary, the behavior in treatment PI crucially depends on whether cooperators always disclose their actions. The more likely they are to disclose, the smaller the difference of action choices will be in RS and PI.

**Hypothesis 2.** *In treatment PI, a higher disclosure rate of those choosing  $x$  leads to a smaller difference of action choices between PI and RS.*

## 5 Experimental Results

### 5.1 Treatment level performances

In the experiment, the number of repetition rounds varies across the eight groups within each treatment, with each group respectively participating in 17, 27, 30, 32, 39, 40, 43, or

47 rounds. To compare participants' behaviors across treatments, we divide the repeated rounds of each group into segments of 10 rounds. Specifically, we analyze participants' behaviors within the following round intervals: 1-10, 11-20, 21-30, 31-40, and 41-50.

Figure 1 presents the evolution of the average cooperation level (represented by the number of players choosing  $x$  within each group) and the average number of neighbors across the three treatments. In the top panel, a notable trend is observed in the RS treatment, wherein most groups demonstrate a consistent increase in the average cooperation level, ultimately converging to the highest level of eight after round 30. Conversely, groups in the CP treatment exhibit fluctuating cooperation levels at a relatively low level without any clear convergence trend. The PI treatment exhibits two distinct evolutionary patterns. In groups 1, 2, 6, and 8 (referred to as PI-C), cooperation levels converge to the maximum level, resembling the behavior observed in the RS treatment. Conversely, the cooperation levels for the remaining four groups (3, 4, 5, and 7), referred to as PI-NC, fluctuate at a relatively low level, resembling the pattern observed in the CP treatment.

The bottom panel illustrates a positive correlation between the dynamics of neighborhood formation and cooperation levels in RS and PI treatments. In RS and PI-C groups, participants' average numbers of neighbors gradually increase over time, eventually reaching the highest level of seven. In contrast, the PI-NC groups maintain a relatively low number of neighbors throughout the experiment. However, the dynamics of neighborhood formation in the CP treatment does not exhibit a significant correlation with cooperation levels. Notably, participants in the CP treatment quickly converge to the complete networks from the initial rounds of the experiment.

To provide statistical support for these observations, Table 1 presents groups' average cooperation level, the average number of neighbors for each participant, and the results of pairwise Mann-Whitney tests across treatments for rounds 1-10, 11-20, and 21-30, respectively.<sup>23</sup> In the CP treatment, the average cooperation level exhibits a declining trend over time, decreasing from 3.96 in rounds 1-10 to 3.34 in rounds 11-20, and further to 2.87 in rounds 21-30. This decreasing pattern of cooperation level aligns with well-established findings in previous studies on public goods or bads games (Isaac

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<sup>23</sup>Table A.1 in Online Appendix D extends Table 1 by presenting the average cooperation level and number of neighbors for rounds 21-40 and 21-50. The averages and significance test results remain largely unchanged compared with rounds 21-30.

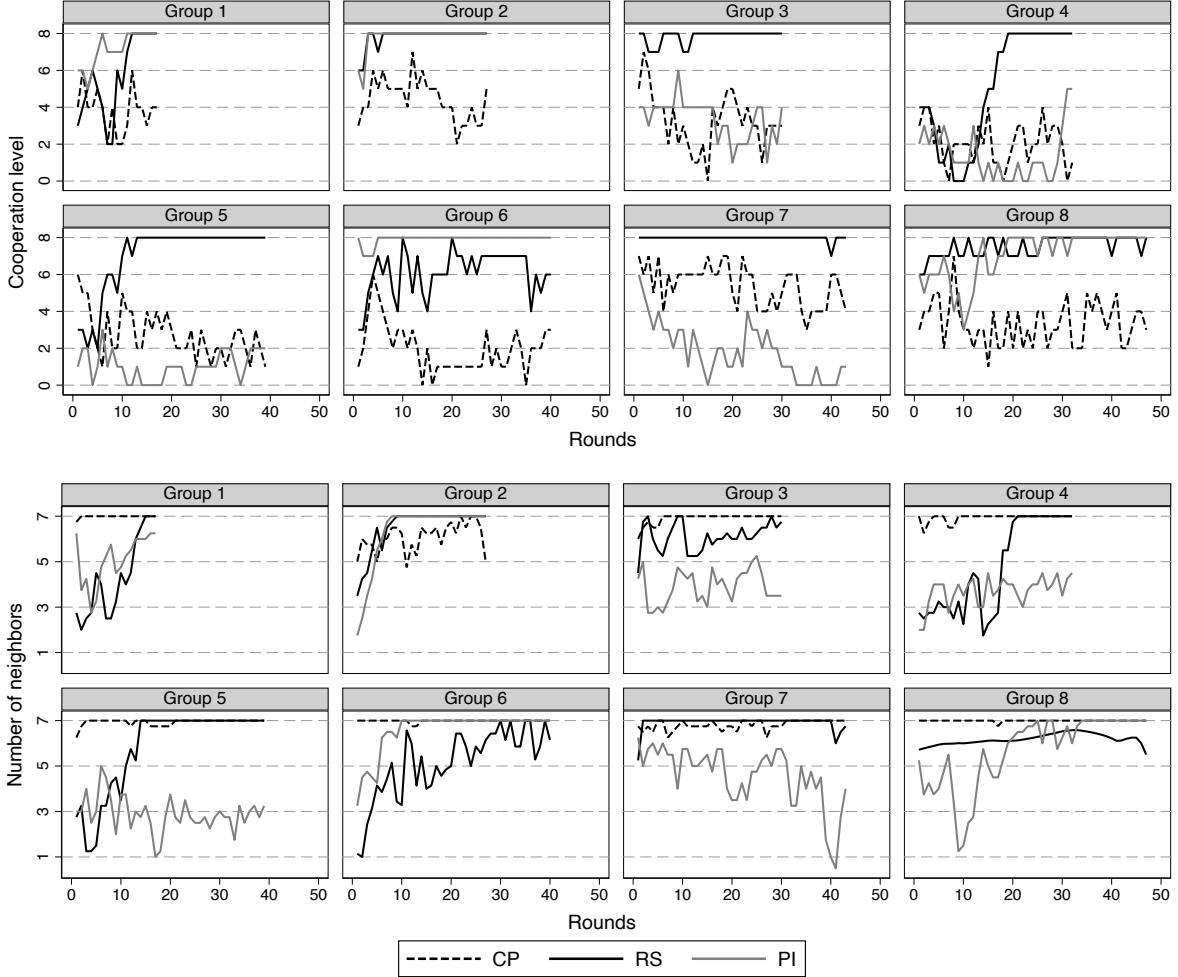


Figure 1: Evolution of groups' cooperation level (top panel) and the number of neighbors (bottom panel).

and Walker, 1988a,b; Ledyard, 1995; Chaudhuri, 2011). However, the cooperation levels in the CP treatment do not converge to nearly zero, as has been commonly observed in studies of standard public goods or bads without punishment (van der Heijden and Moxnes, 1999; Fehr and Gächter, 2000; Moxnes and Van der Heijden, 2003; Gurerk et al., 2006; Lugovskyy et al., 2017). This discrepancy could potentially be attributed to two distinctive features of our setting: (i) the presence of centralized punishment, and (ii) the network formation stage requiring mutual consent. We investigate this channel in more detail in Section 5.2.2. In contrast, the average cooperation level in the RS treatment exhibits an increasing trend over time, rising from 5.71 in rounds 1-10 to 7.23 in rounds 11-20, and further to 7.77 in rounds 21-30. These levels are significantly higher than those in the CP treatment, as indicated by Mann-Whitney tests (rounds 1-10,  $p = 0.047$ ;

Table 1: Statistics of cooperation levels and number of neighbors

| Rounds                               | Average cooperation level |        |         | Average number of neighbors |        |        |
|--------------------------------------|---------------------------|--------|---------|-----------------------------|--------|--------|
|                                      | Maximal level: 8          |        |         | Maximal number: 7           |        |        |
| 1-10                                 | 11-20                     | 21-30  | 1-10    | 11-20                       | 21-30  |        |
| CP                                   | 3.96                      | 3.34   | 2.87    | 6.75                        | 6.82   | 6.90   |
|                                      | (1.07)                    | (1.64) | (1.13)  | (0.37)                      | (0.33) | (0.21) |
| RS                                   | 5.71                      | 7.23   | 7.77    | 4.57                        | 5.91   | 6.69   |
|                                      | (2.16)                    | (1.19) | (0.45)  | (1.72)                      | (1.03) | (0.40) |
| PI                                   | 4.79                      | 4.54   | 4.36    | 4.34                        | 4.99   | 5.16   |
|                                      | (2.41)                    | (3.47) | (3.42)  | (0.91)                      | (1.57) | (1.70) |
| Mann-Whitney test ( <i>p</i> -value) |                           |        |         |                             |        |        |
| CP vs. RS                            | 0.047                     | 0.001  | < 0.001 | 0.005                       | 0.078  | 0.339  |
| CP vs. PI                            | 0.505                     | 0.591  | 0.870   | < 0.001                     | 0.047  | 0.054  |
| RS vs. PI                            | 0.367                     | 0.359  | 0.078   | 1.000                       | 0.218  | 0.143  |

Notes: Standard deviations are in parentheses. Two-sided Mann-Whitney tests are performed at the group level ( $n = 16$  for rounds 1-10 and rounds 11-20;  $n = 14$  for rounds 21-30). Only seven groups remain in rounds 21-30 of each treatment as one group ends at round 17.

rounds 11-20,  $p = 0.001$ ; rounds 21-30,  $p < 0.001$ ). In comparison, the PI treatment show no clear trend in either cooperation levels or number of neighbors. The cooperation level in PI is slightly lower than in RS for rounds 21-30 ( $p = 0.078$ ).

The statistics of participants' number of neighbors also support the observations depicted in Figure 1. In the CP treatment, almost all participants quickly form a complete network. In contrast, participants in the RS treatment initially establish fewer neighbors than in CP treatment (rounds 1-10,  $p = 0.005$ ; rounds 11-20,  $p = 0.078$ ). However, over time the number of neighbors steadily rises in RS, reaching nearly seven in rounds 21-30, which is statistically indistinguishable from CP ( $p = 0.339$ ). In the PI treatment, the number of neighbors is marginally lower than in RS across all three segments of rounds, although the difference is consistently insignificant.

In summary, these findings demonstrate the substantial impact of the RS mechanism in inhibiting public bad actions when choices are publicly observable, supporting Hypothesis

1. However, when actions remain private, the RS mechanism fosters cooperation in only half of the groups and does not yield significantly higher overall cooperation than CP.

**Result 1.** *Participants in CP quickly form complete network, but cooperation decreases over time while remaining above zero. Consistent with Hypothesis 1, the RS mechanism significantly reduces public bad actions, increases network size, and improves cooperation levels. In PI, outcomes diverge, with about half the groups converging to full cooperation as in RS, while the remaining groups do not.*

## 5.2 Determinants of individual-level choices

In this section, we examine the determinants of participants' link decisions and public bad choices across the three treatments. We investigate participants' information disclosure behavior in the PI treatment.

### 5.2.1 Determinants of link decisions

First, we present the results of regressions using the linear probability model (LPM) with individual and round fixed effects to examine the determinants of link decisions in Panel A of Table 2.<sup>24</sup> Columns (1)–(3) correspond to each of the three treatments and columns (4)–(5) provide a breakdown of the PI treatment into four successful cooperation groups (PI-C) and the remaining four groups (PI-NC). The dummy dependent variables represent arbitrary participant  $i$ 's link decisions to arbitrary participant  $j$  ( $j \in N, j \neq i$ ). A value of 1 indicates that participant  $i$  proposes the link, while 0 indicates no proposal. The first five explanatory variables represent the joint actions taken by participants  $i$  and  $j$  in round  $t - 1$ . Specifically, dummy variables  $x_{i,t-1}$  and  $y_{i,t-1}$  respectively equal 1 if participant  $i$  chooses action  $x$  and  $y$ , and  $x_{j,t-1}$ ,  $y_{j,t-1}$ , and  $nd_{j,t-1}$  equal 1 if participant  $j$  is observed choosing action  $x$ ,  $y$ , and does not disclose the action (only in PI), respectively.<sup>25</sup> Note that in columns (3)–(5), if participant  $j$ 's action can be inferred as choosing action  $x$  or  $y$ ,

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<sup>24</sup>Two fixed-effects dimensions arise in link decisions, the link proposer  $i$  and the round  $t$ . We employ the LPM approach, which flexibly and efficiently accommodates high-dimensional fixed effects while yielding coefficients with a straightforward and intuitive interpretation.

<sup>25</sup>In columns (1) and (2) of Table 2, the benchmark case (omitted in the regressions) corresponds to the case in which both participants  $i$  and  $j$  choose action  $y$ . In columns (3) to (5), the benchmark case represents the scenario in which participant  $i$  chooses action  $y$  and participant  $j$  does not disclose their action.

we let  $x_{j,t-1} = 1$  or  $y_{j,t-1} = 1$ .<sup>26</sup> Next, the explanatory variable,  $neighbor_{ij,t-1}$  denotes the link history between participants  $i$  and  $j$ , equaling 1 if  $i$  and  $j$  were neighbors in round  $t - 1$  and 0 otherwise. Finally, the explanatory variable,  $reputation_{j,1 \sim t-2}$  records player  $j$ 's rate of observable or inferrable  $x$  choices in rounds 1 to  $t - 2$ . In treatments CP and RS, this rate is simply the rate of  $x$  choices, whereas in the treatment PI, choice  $x$  is observed in the case of disclosing or inferred in the case of group-level full cooperation. To facilitate cross-treatment interpretation, Panel B of Table 2 reports Wald tests of coefficient equality across treatments for the key comparisons discussed below.

We first examine whether participants' tendencies to link with others differ based on their action choices in the previous rounds. The significance patterns of variables  $\beta_1$ ,  $\beta_2$ ,  $\beta_3$ , and  $\beta_8$  in columns (1) and (2) of Table 2 indicate that participants in the CP treatment who chose action  $x$  in the previous round (cooperative participants) and those who chose action  $y$  in the previous round (noncooperative participants) exhibit no significant preferences in link decisions toward cooperative or noncooperative others, nor with respect to others' reputations. In other words, participant's link decisions in the CP treatment are independent of others' past behavior. By contrast, all participants in the RS treatment, particularly cooperative ones, propose significantly more links to other cooperative participants than to noncooperative participants and the number of proposals also significantly increases with the reputation levels of others.<sup>27</sup> This indicates that cooperative participants in the RS treatment adopt a *link reciprocity* strategy, making strategic proposal decisions not only based on behavior in the previous round but also on behavior in earlier rounds. This finding is intuitive: due to the responsibility-sharing feature, participants in RS rely on all available information to assess the potential risks of forming links with others. When this concern is absent in CP, participants need not consider others' past behavior since linking with them only brings benefits.

In the PI treatment, when action choices are not disclosed,  $\beta_2$ ,  $\beta_4$ , and  $\beta_5$  are not significantly different from zero, indicating that participants exhibit similar linking tendencies toward noncooperative and nondisclosure participants. Moreover, separately

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<sup>26</sup>Specifically, a few scenarios emerge in which participant  $j$ 's action can be inferred: (i) when they disclose their actions, (ii) when the feedback information indicates that all others (including participant  $j$ ) choose  $x$  or  $y$ , and (iii) participant  $j$  is detected to have chosen action  $y$ .

<sup>27</sup>These cross-treatment differences between CP and RS are confirmed by the Wald tests reported in Panel B of Table 2.

Table 2: Determinants of participant  $i$ 's link decisions in each treatment

| <i>Panel A. Treatment-specific regressions</i> |                       |                       |                      |                      |                      |
|--|-----------------------|-----------------------|----------------------|----------------------|----------------------|
|  | (1)<br>CP             | (2)<br>RS             | (3)<br>PI            | (4)<br>PI-C          | (5)<br>PI-NC         |
| $\beta_1 : x_{i,t-1} \times x_{j,t-1}$         | 0.00604<br>(0.00335)  | 0.222***<br>(0.0314)  | 0.226**<br>(0.0706)  | 0.362***<br>(0.0235) | 0.0520<br>(0.0451)   |
| $\beta_2 : x_{i,t-1} \times y_{j,t-1}$         | -0.00864<br>(0.00941) | -0.183***<br>(0.0493) | -0.159<br>(0.0904)   | -0.320**<br>(0.0624) | -0.113<br>(0.101)    |
| $\beta_3 : y_{i,t-1} \times x_{j,t-1}$         | 0.0106<br>(0.00878)   | 0.195***<br>(0.0232)  | 0.119*<br>(0.0585)   | 0.336***<br>(0.0383) | 0.0137<br>(0.0247)   |
| $\beta_4 : y_{i,t-1} \times y_{j,t-1}$         |                       | (benchmark)           |                      | -0.0670<br>(0.0370)  | -0.156<br>(0.0745)   |
| $\beta_5 : x_{i,t-1} \times nd_{j,t-1}$        |                       |                       |                      | -0.0199<br>(0.0335)  | 0.0221<br>(0.0600)   |
| $\beta_6 : y_{i,t-1} \times nd_{j,t-1}$        |                       |                       |                      |                      | (benchmark)          |
| $\beta_7 : neighbor_{ij,t-1}$                  | 0.158**<br>(0.0471)   | 0.175***<br>(0.0334)  | 0.262***<br>(0.0355) | 0.240***<br>(0.0281) | 0.261**<br>(0.0516)  |
| $\beta_8 : reputation_{j,1 \sim t-2}$          | 0.0124<br>(0.00779)   | 0.213**<br>(0.0641)   | 0.0977**<br>(0.0333) | 0.0788*<br>(0.0319)  | 0.173*<br>(0.0656)   |
| constant                                       | 0.827***<br>(0.0482)  | 0.422***<br>(0.0636)  | 0.516***<br>(0.0218) | 0.342***<br>(0.0394) | 0.580***<br>(0.0269) |
| <i>N</i>                                       | 14504                 | 14504                 | 14504                | 6888                 | 7616                 |

| <i>Panel B. Key cross-treatment comparisons (Wald tests)</i> |  |                 |
|--|--|-----------------|
| Variables  | Null hypothesis  | <i>p</i> -value |
| $\beta_1 : x_{i,t-1} \times x_{j,t-1}$                       | $\beta_1^{CP} = \beta_1^{RS}$  | < 0.001         |
| $\beta_2 : x_{i,t-1} \times y_{j,t-1}$                       | $\beta_2^{CP} = \beta_2^{RS}$  | = 0.002         |
| $\beta_3 : y_{i,t-1} \times x_{j,t-1}$                       | $\beta_3^{CP} = \beta_3^{RS}$  | < 0.001         |
| $\beta_8 : reputation_{j,1 \sim t-2}$                        | $\beta_8^{CP} = \beta_8^{RS}$  | = 0.003         |
| Joint test   | $\{\beta_{1 \sim 5,7,8}^{PI-C}\} = \{\beta_{1 \sim 5,7,8}^{PI-NC}\}$ | < 0.001         |
| Joint test   | $\{\beta_{1 \sim 3,7,8}^{CP}\} = \{\beta_{1 \sim 3,7,8}^{PI-NC}\}$   | < 0.001         |
| Joint test   | $\{\beta_{1 \sim 3,7,8}^{RS}\} = \{\beta_{1 \sim 3,7,8}^{PI-C}\}$    | < 0.001         |

Standard errors (in parentheses) are clustered at the group level.

Panel B reports Wald tests of coefficient equality across treatments based on pooled specifications with treatment interactions, using the same controls and fixed effects as in Panel A ( $N = 43512$ ).

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

analyzing the PI-C and PI-NC groups reveals that participants in the two groups behave differently in their linking decisions (Wald joint test:  $p < 0.001$ ). Specifically, the similar signs and statistical significance of the coefficients suggest that participants in PI-C behave qualitatively similarly to those in the RS treatment, whereas participants in PI-NC resemble those in the CP treatment.<sup>28</sup>

These findings indicate that the link reciprocity strategy is rarely adopted in our global public bad games with centralized punishment, in contrast to findings from local public goods/bads games (Cinyabuguma et al., 2005; Page et al., 2005; Rand et al., 2011; Charness and Yang, 2014). This is arguably because linking with others in our CP setting yields only benefits and no costs. By contrast, becoming neighbors with individuals who may choose the public bad action under the RS mechanism can result in potential loss. Therefore, participants must weigh the benefits of any link against its expected costs, making link reciprocity a natural strategy. Furthermore, the distinct linking strategies employed by PI-C and PI-NC participants may help explain overall cooperation disparity between these two groups. We revisit this point after discussing participants' public bad actions and disclosure behavior.

**Result 2.** *In RS and PI-C, cooperative and noncooperative participants are more likely to link with cooperators than non-cooperators, whereas in CP and PI-NC they show no such tendency. In the PI treatment, participants do not differentiate between non-disclosure and noncooperative participants when forming links.*

### 5.2.2 Determinants of public bad actions

Next, we analyze the factors influencing participants' decisions to choose public bad action  $y$ . Panel A of Table 3 presents the marginal effects from random-effects probit regressions, examining participant  $i$ 's choice of action  $y$  in round  $t$  for each treatment.<sup>29</sup> The dependent variable is a dummy that equals 1 if participant  $i$  chooses action  $y$  in round  $t$ . The explanatory variables include lagged variables from round  $t - 1$  such as whether participant  $i$  chose action  $y$  ( $y_{i,t-1}$ ), other participants' cooperation rate in the group ( $\%cooperation_{j \neq i, t-1}$ ), whether participant  $i$  was detected regardless of

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<sup>28</sup>However, the Wald joint tests reject the equality of coefficients in both cases ( $p < 0.001$ ).

<sup>29</sup>Table A.2 in Online Appendix D presents the results of fixed-effects logit regressions, demonstrating the robustness of Table 3.

their choice ( $detection_{i,t-1}$ ), whether they were punished for being detected choosing  $y$  ( $punished_{i,t-1}^{detection}$ ), and whether they were punished owing to their neighbors' behavior ( $punished_{i,t-1}^{neighbor}$ ). Additional explanatory variables include the number of neighbors in round  $t$  ( $\#neighbors_{i,t}$ ), the number of failed proposals ( $\#failed_{i,t}$ ) in round  $t$ , and its interaction with  $y_{i,t-1}$ . Panel B of Table 3 reports Wald tests of coefficient equality across treatments for the key comparisons discussed below.

The regression results demonstrate the significant influence of other participants' cooperation rate ( $\beta_2$ ) on an individual's decision to refrain from taking public bad actions across all three treatments. This finding supports the existence of "conditional cooperators," who are more likely to cooperate when they observe others doing so (Keser and Van Winden, 2000; Fischbacher et al., 2001; Kurzban and Houser, 2005; Rustagi et al., 2010; Fischbacher and Gächter, 2010; Yang et al., 2018; Chen et al., 2026). Moreover, while cooperative and noncooperative participants in the RS treatment in Table 2 exhibit a significant aversion to establishing links with noncooperators, the number of failed proposals ( $\beta_4$ ) does not directly affect participants' choices of public bad actions. Instead, the number of neighbors ( $\beta_3$ ) emerges as a significant factor that discourages participants from choosing action  $y$  in the RS treatment.

By combining the determinants of participants' link decisions and public bad choices, we conclude that successful cooperation in the RS treatment is attributable to participants' preference for linking with cooperators and avoiding links with noncooperators. This exclusionary behavior deters non-cooperators from choosing the public bad action, particularly when they form more neighbors, as they fear future exclusion and the resulting loss of network benefits. However, when participants do not distinguish between cooperators and noncooperators in link formation, noncooperators no longer face this threat, and refraining from the public bad action is no longer advantageous.

In the CP treatment, apart from the cooperation rate in the previous round ( $\beta_2$ ), the only factor that discourages noncooperative participants from choosing action  $y$  is their number of failed proposals ( $\beta_5$ ). This indicates that noncooperative participants are more likely to switch to cooperation when they experience more failed proposals. Moreover, the larger absolute value of  $\beta_4$  compared with  $\beta_5$  indicates that when cooperative participants' proposals fail, they are even more likely to switch to noncooperation. In other words, both cooperative and noncooperative participants adjust their behavior in

Table 3: Determinants of participant  $i$ 's public bad actions in each treatment

| Panel A. Treatment-specific regressions     |                      |                      |                      |                      |                      |
|---|----------------------|----------------------|----------------------|----------------------|----------------------|
|   | (1)<br>CP            | (2)<br>RS            | (3)<br>PI            | (4)<br>PI-C          | (5)<br>PI-NC         |
| $\beta_1 : y_{i,t-1}$                       | 0.154***<br>(0.037)  | 0.091***<br>(0.031)  | 0.228***<br>(0.019)  | 0.053***<br>(0.019)  | 0.266***<br>(0.038)  |
| $\beta_2 : \%cooperation_{j \neq i, t-1}$   | -0.034***<br>(0.010) | -0.029***<br>(0.006) | -0.043***<br>(0.006) | -0.006***<br>(0.002) | -0.030**<br>(0.013)  |
| $\beta_3 : \#neighbors_{i,t}$               | 0.100***<br>(0.024)  | -0.015**<br>(0.007)  | -0.030***<br>(0.003) | -0.019***<br>(0.004) | -0.033***<br>(0.004) |
| $\beta_4 : \#failed_{i,t}$                  | 0.323***<br>(0.071)  | 0.005<br>(0.009)     | 0.003<br>(0.010)     | 0.000<br>(0.005)     | -0.002<br>(0.017)    |
| $\beta_5 : y_{i,t-1} \times \#failed_{i,t}$ | -0.228***<br>(0.056) | -0.010<br>(0.008)    | -0.009<br>(0.008)    | -0.007<br>(0.008)    | 0.006<br>(0.010)     |
| $\beta_6 : detection_{i,t-1}$               | 0.229***<br>(0.063)  | 0.045*<br>(0.024)    | -0.040<br>(0.035)    | 0.004<br>(0.009)     | -0.121**<br>(0.059)  |
| $\beta_7 : punished_{i,t-1}^{detection}$    | -0.086<br>(0.067)    | 0.007<br>(0.025)     | -0.037<br>(0.035)    | 0.038<br>(0.033)     | -0.027<br>(0.068)    |
| $\beta_8 : punished_{i,t-1}^{neighbor}$     |                      | -0.009<br>(0.020)    | 0.012<br>(0.028)     | 0.016<br>(0.011)     | 0.004<br>(0.031)     |
| $N$   | 2136                 | 2136                 | 2136                 | 1016                 | 1120                 |

| Panel B. Key cross-treatment comparisons (Wald tests on marginal effects) |  |            |
|---|--|------------|
| Variables   | Null hypothesis  | $p$ -value |
| $\beta_3 : \#neighbors_{i,t}$   | $\beta_3^{CP} = \beta_3^{RS}$                                | = 0.006    |
| Joint test  | $\{\beta_{1 \sim 8}^{RS}\} = \{\beta_{1 \sim 8}^{PI}\}$      | < 0.001    |
| Joint test  | $\{\beta_{1 \sim 8}^{PI-C}\} = \{\beta_{1 \sim 8}^{PI-NC}\}$ | < 0.001    |

Standard errors (in parentheses) are clustered at the group level.

Panel B reports Wald tests of equality of treatment-specific average marginal effects. The tests are implemented using post-estimation contrasts based on margins and the delta method, with the same clustering level as in Panel A ( $N = 6408$ ).

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

response to failed proposals, preventing behavioral convergence. The intuition is that, because mutual consent is required to establish neighbors, failed proposals discourage cooperators from continuing to cooperate, while motivating noncooperative participants to start cooperating instead. This switching pattern explains why the the CP treatment' cooperation rate fluctuates over time rather than converging to zero, as is commonly observed in the literature on standard public good games. Furthermore, being punished by the centralized punishment itself does not significantly affect behavior in the CP treatment ( $\beta_7$ ). However, it is possible that participants already consider the probability of being punished when making their decisions, regardless of whether they are actually punished. In summary, we demonstrate that the mutual consent requirement in network formation can explain why behavior fails to converge in the CP treatment, but no direct evidence for the centralized punishment mechanism is found.

In contrast to the RS treatment, the number of neighbors ( $\beta_3$ ) in the CP treatment does not discourage participants from choosing the public bad (Wald test,  $\beta_3^{RS} = \beta_3^{CP} : p = 0.006$ ). This finding aligns with the conclusion that neither cooperative nor noncooperative participants in CP show a significant aversion to linking with non-cooperators, as shown in Table 2. Consequently, no mechanism to deter noncooperative participants from choosing action  $y$  is evident, which may explain the lower cooperation level observed in CP. Notably, when a participant is detected in round  $t-1$  in both CP and RS treatments, it encourages them to choose the public bad in round  $t$  ( $\beta_6$ ). This implies that being detected *per se* may backfire in our setting. Finally, receiving punishment, regardless of its sources, has no significant effect on cooperation behavior.

In the PI treatment, the determinants of public bad choices exhibit qualitatively similar patterns (e.g., similar signs and statistical significance of the coefficients) to those in the RS treatment and across PI-C and PI-NC groups, although the coefficients differ statistically.<sup>30</sup>

**Result 3.** *Participants exhibit conditional cooperation in all three treatments. In RS and PI treatments, a larger number of neighbors significantly reduces the likelihood of choosing public bad actions, whereas the opposite holds in the CP treatment. The determinants of public bad actions exhibit similar qualitative patterns between PI-C and PI-NC, although the estimated effects differ statistically across the two groups.*

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<sup>30</sup>The Wald joint tests on marginal effects reject equality in both cases ( $p < 0.001$ ).

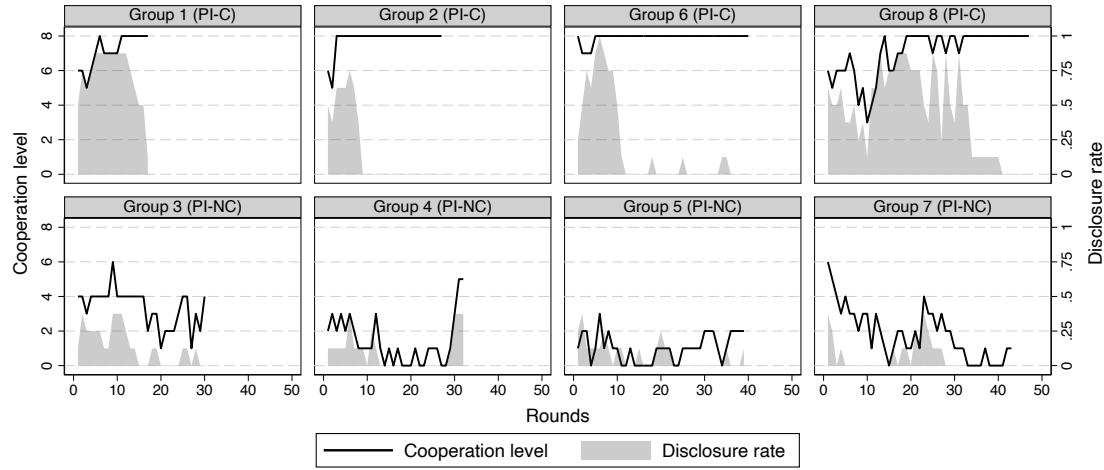


Figure 2: Cooperation level and disclosure rate evolution in PI

### 5.2.3 Information disclosure in PI

This section analyzes the dynamics and determinants of information disclosure behavior in the PI treatment and explains PI-C and PI-NC groups' heterogeneous performance. Figure 2 presents each group's disclosure rate evolution alongside cooperation levels for comparison. To enhance clarity, we present PI-C groups in the top panel and PI-NC groups in the bottom panel. The figure reveals a notable disparity in information disclosure between PI-C and PI-NC before convergence. Specifically, in rounds 1-10, participants in PI-C demonstrate a significantly higher disclosure rate of 59% compared with 16% in PI-NC (Mann-Whitney test,  $p = 0.029$ ,  $n = 8$ ). However, once full cooperation is reached, the disclosure rate in PI-C sharply declines to 20% in rounds 21-30, at which point it no longer differs significantly from PI-NC (Mann-Whitney test,  $p = 0.629$ ,  $n = 7$ ). In rounds where groups have not yet achieved full cooperation, cooperative participants' disclosure rates are significantly higher in PI-C compared with PI-NC (73.0% versus 31.1%; Mann-Whitney test,  $p = 0.029$ ,  $n = 8$ ). Finally, it is noteworthy that 95.4% of all disclosed decisions in the PI treatment were  $x$  choices, indicating that the primary purpose of disclosure is reputation formation.

Panel A of Table 4 presents the marginal effects from random-effects probit regressions that investigate the determinants of participant  $i$ 's decisions to disclose action  $x$  in PI, PI-C, and PI-NC groups.<sup>31</sup> The dependent variable is a binary indicator that equals

<sup>31</sup>We specifically analyze disclosure behavior when participants choose to disclose action  $x$ , as it represents the majority of disclosed information and is closely correlated with the cooperation success

1 if participant  $i$  discloses action  $x$  in round  $t$ . The regressions examine the impact of several explanatory variables, encompassing whether participant  $i$  disclosed action  $x$  in the previous round ( $d_{i,t-1}$ ), the group's disclosure rate of  $x$  choices in the previous round ( $\%d_{t-1}$ ), the number of participant  $i$ 's failed proposals in the current round ( $\#failed_{i,t}$ ). We also introduce a set of dummy variables to capture different cooperation levels in the current round, divided into low ( $lowcoop_t$ , 0–3 players choose  $x$ ), high ( $highcoop_t$ , 4–7 players choose  $x$ ), and full ( $fullcoop_t$ , all eight players choose  $x$ ) cooperation. Further explanatory variables include whether participant  $i$  chose action  $y$  in the previous round ( $y_{i,t-1}$ ) and participant  $i$ 's number of neighbors in the current round ( $\#neighbors_{i,t}$ ). Panel B of Table 4 reports Wald tests of coefficient equality between PI-C and PI-NC for the key comparisons discussed below.

Column (1) of Panel A shows that individual's disclosure behavior, the rest of the group's disclosure rate in the previous round, the number of failed proposals, and the number of neighbors in the current round all have a significant positive effect on disclosure. The effect of cooperation levels on disclosing exhibits an inverted U-shape, wherein compared with the high cooperation benchmark, the disclosure rate is significantly lower in both low- ( $\beta_5$ ) and full-cooperation ( $\beta_7$ ) cases.<sup>32</sup> This finding indicates that as cooperation increases, participants initially become more likely to disclose their  $x$  choices; however, they reduce disclosure once full cooperation is achieved. Columns (2) and (3) of Panel A reveal that the positive effects of the past group disclosure rate and the number of neighbors are significant only in PI-NC, while the positive effect of failed proposals is stronger in PI-C than in PI-NC (Wald test on marginal effects:  $p = 0.030$ ).

Thus far, we have identified two key factors in participants' decision making that distinguish the PI-C and PI-NC groups. First, results in Section 5.2.1 indicate that participants in PI-C exhibit preferences similar to those in the RS treatment when forming connections, whereas those in PI-NC display preferences resembling those in the CP observed in the PI-C group. For a comprehensive examination of the determinants of disclosure behavior irrespective of chosen actions, see Table A.3 in Online Appendix D. In addition, Table A.4 in Online Appendix D presents the results of fixed-effects logit regressions, which confirm the robustness of Table 4.

<sup>32</sup>Because the effect of cooperation levels is non-monotonic, we employ three dummy variables representing different cooperation levels, rather than using a single explanatory variable for cooperation level.

Table 4: Determinants of participant  $i$ 's disclosure of action  $x$ 

| <i>Panel A. Treatment-specific regressions</i> |                      |                      |                      |
|--|----------------------|----------------------|----------------------|
|  | (1)<br>PI            | (2)<br>PI-C          | (3)<br>PI-NC         |
| $\beta_1 : d_{i,t-1}$                          | 0.108***<br>(0.021)  | 0.129***<br>(0.037)  | 0.108***<br>(0.033)  |
| $\beta_2 : \%d_{t-1}$                          | 0.179**<br>(0.070)   | 0.163<br>(0.132)     | 0.102***<br>(0.005)  |
| $\beta_3 : \#failed_{i,t}$                     | 0.028***<br>(0.006)  | 0.033***<br>(0.008)  | 0.017**<br>(0.008)   |
| $\beta_4 : \#neighbors_{i,t}$                  | 0.017***<br>(0.005)  | 0.021<br>(0.017)     | 0.015***<br>(0.005)  |
| $\beta_5 : lowcoop_t$                          | -0.119**<br>(0.047)  | -0.220***<br>(0.027) | -0.052***<br>(0.007) |
| $\beta_6 : highcoop_t$                         |                      | (benchmark)          |                      |
| $\beta_7 : fullcoop_t$                         | -0.229***<br>(0.043) | -0.308***<br>(0.066) |                      |
| $\beta_8 : y_{i,t-1}$                          | -0.032<br>(0.020)    | 0.014<br>(0.071)     | -0.013<br>(0.013)    |
| <i>N</i>                                       | 2136                 | 1016                 | 1120                 |

| <i>Panel B. Key cross-treatment comparisons (Wald tests on marginal effects)</i> |  |                 |
|--|--|-----------------|
| Variables  | Null hypothesis  | <i>p</i> -value |
| $\beta_3 : \#failed_{i,t}$   | $\beta_3^{PI-C} = \beta_3^{PI-NC}$                               | = 0.030         |
| Joint test   | $\{\beta_{1 \sim 5,8}^{PI-C}\} = \{\beta_{1 \sim 5,8}^{PI-NC}\}$ | < 0.001         |

Standard errors (in parentheses) are clustered at the group level.

Panel B reports Wald tests of equality of treatment-specific average marginal effects. The tests are implemented using post-estimation contrasts based on margins and the delta method, with the same clustering level as in Panel A ( $N = 2136$ ).

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

treatment. Second, cooperative participants' disclosure rates are higher in PI-C than in PI-NC. Combining these findings can potentially explain the higher level of cooperation observed in the PI-C group. Cooperative participants' increased disclosure rate makes the decision-making environment in PI-C similar to that of the RS treatment. As a result, cooperative participants are more likely to establish stable and extensive networks with one another, motivating noncooperators to connect with cooperators, refrain from choosing the public bad action, and thereby foster successful cooperation within the PI-C groups.

In summary, our findings reveal the significance of cooperative participants' early information disclosure in facilitating successful cooperation within the PI treatment. Early disclosure enables cooperators to connect with one another and encourages noncooperators to adopt cooperative behavior. However, once full cooperation is achieved, the role of disclosure is diminished, as participants can sustain cooperation without relying heavily on it.

**Result 4.** *Cooperative participants' disclosure rate is higher in PI-C than in PI-NC, supporting Hypothesis 2. Information disclosure is essential before cooperation converges to the highest level but becomes unnecessary once full cooperation is achieved.*

### 5.3 Efficiency

This section analyzes efficiency loss across treatments, defined as the reduction in participants' payoffs relative to the full-cooperation outcome. While the RS mechanism effectively discourages public bad actions, it also introduces additional costs by forming fewer links, particularly in groups and rounds that have not yet reached full cooperation. In addition, responsibility sharing may generate further costs through punishment. Accordingly, we decompose participants' efficiency loss into three components for rounds 1-10, 11-20, and 21-30, as shown in Figure 3. The first component captures the loss from choosing public bad actions.<sup>33</sup> The second component measures scale loss (i.e., reduction from forming fewer neighbors than the efficient complete network. The third component represents the total punishment imposed on participants.

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<sup>33</sup>This component has two parts. The first is the cost of negative externalities caused by others' public bad actions. The second is the gain from individual's choosing the public bad action. Therefore, we define the loss from public bad actions as the net difference between these two parts.

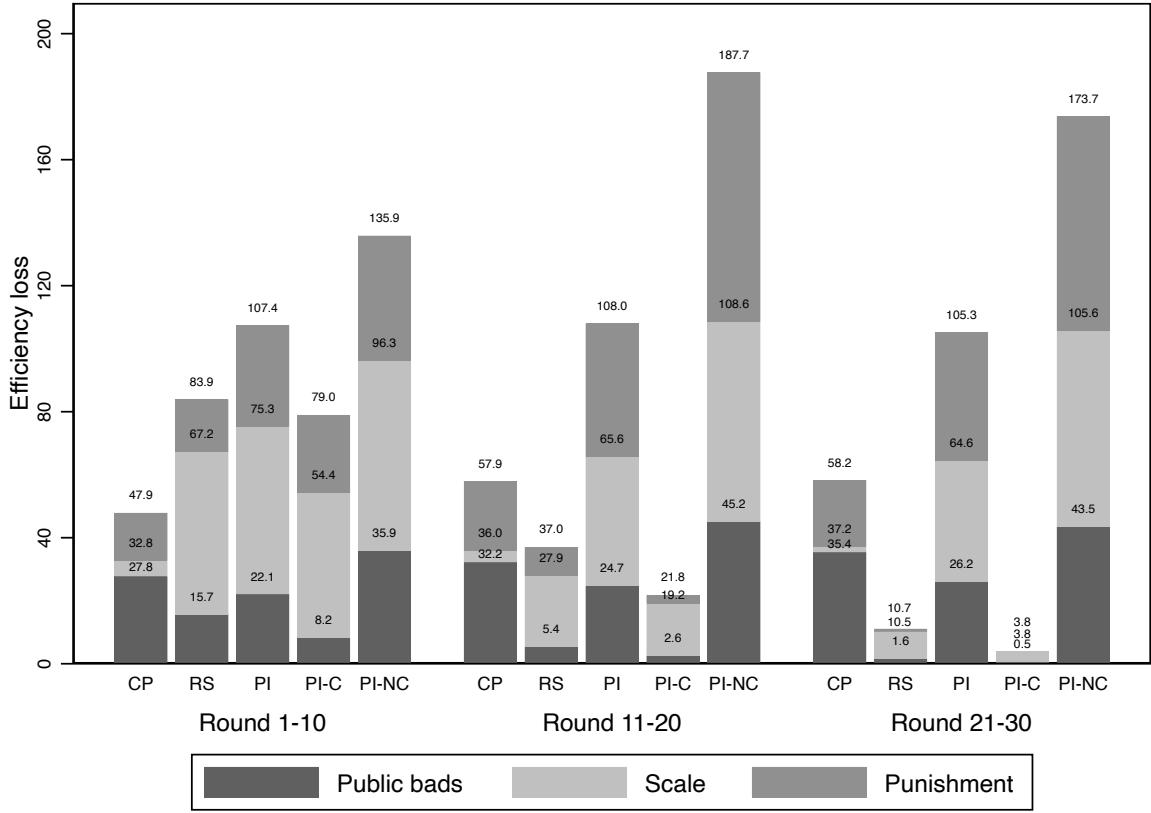


Figure 3: **Decomposition of average efficiency losses for each treatment** (the level of accumulated loss is presented above each bar).

Efficiency losses in the CP treatment remain steady over time, primarily driven by frequent public bad choices and resulting punishments. In the first 10 rounds of the RS treatment, most efficiency loss comes from the scale component, as participants do not form large networks during these rounds. Notably, the loss from public bad actions is significantly lower in RS than in CP (Mann-Whitney test,  $p = 0.047$ ,  $n = 16$ ). Moreover, although participants share punishment in RS, the efficiency loss from punishments does not differ significantly from that of CP (Mann-Whitney test,  $p = 0.878$ ,  $n = 16$ ). After round 10, efficiency losses from all three components in RS gradually decline, and by rounds 21-30, become significantly lower than in CP (Mann-Whitney test,  $p = 0.001$ ,  $n = 14$ ).

In the PI treatment, overall efficiency loss across the three segments remains high and does not change over time. Dividing the PI treatment into PI-C and PI-NC groups reveals a contrasting pattern. In PI-C, where the RS mechanism successfully promotes cooperation, efficiency loss gradually declines toward zero. In contrast, in PI-NC, where

the mechanism fails to improve cooperation, efficiency loss increases over time. Notably, in rounds 21–30, efficiency loss in PI-NC is even significantly greater than in CP (Mann-Whitney test,  $p = 0.006$ ,  $n = 11$ ). A closer comparison between PI-NC and CP reveals no significant difference in the loss from public bad actions (Mann-Whitney test,  $p = 0.109$ ,  $n = 11$ ). Instead, the greater loss in PI-NC arises mainly from two sources, (i) network scale loss due to fewer neighbors (Mann-Whitney test,  $p = 0.006$ ,  $n = 11$ ), potentially because participants in PI-NC shy away from linking with non-cooperators to avoid responsibility sharing; and (ii) greater punishment imposed under the responsibility-sharing rule (Mann-Whitney test,  $p = 0.006$ ,  $n = 11$ ).

**Result 5.** *Efficiency loss in RS is significantly smaller than in CP. In PI, substantial losses arise when the RS mechanism fails, driven by reduced network size and additional punishments from responsibility sharing.*

## 6 Conclusion

This study investigates the effect of the responsibility-sharing mechanism on controlling public bad actions employing an infinitely repeated two-stage game in an eight-player economy. In each period, players engage in two stages. First, players endogenously form a network and they decide whether to engage in a public bad action that provides individual benefits but harms other players' welfare. A central institution independently and randomly detects players' action choices. Under the RS mechanism, a player is punished if either they or one of their neighbors is detected to have chosen the public bad action.

Theoretical analysis of the game demonstrates that both noncooperation, where all players choose the public bad action, and full cooperation, where all players refrain from it, yield equilibrium outcomes. However, using the basin of attraction criterion, we determine that full cooperation is more likely to emerge under the RS mechanism. We further demonstrate that these results remain robust across a broad range of parameters.

We also conduct an experiment to test the effect of the RS mechanism on mitigating public bad actions, revealing that the mechanism effectively reduces public bad provisions when participants' actions are publicly known. In the treatment without the mechanism (CP treatment), participants tend to form complete networks and frequently choose

the public bad action in each period. In contrast, in the RS treatment, participants tend to avoid linking with non-cooperative players. As a result, noncooperative players refrain from providing the public bad to build links with more players, resulting in full cooperation. In addition, we investigate the effectiveness of the mechanism when players' actions are private information and disclosure is voluntary. Our results indicate that, the mechanism is only partially effective in this setting, with convergence to cooperation depending crucially on the extent of information disclosure.

To the best of our knowledge, this is the first study to investigate the effect and limitations of responsibility sharing among endogenously formed neighbors in the context of public bad. Our research opens several avenues for future inquiry. First, in many real-life scenarios, interactions are formed through centralized institutions rather than decentralized network formation. Further exploration is needed to understand how these institutional differences may affect strategies and choices under RS. Second, while we model RS as equal punishment for all players within a network, other forms of RS—such as alternative punishment allocation rules—can yield different outcomes. Exploring more nuanced design could deepen our understanding of how punishment allocation rules shape cooperation. Finally, regarding information availability, an alternative assumption is that players only observe the actions of their direct neighbors, rather than the entire group as in our experiment. We conjecture that the RS mechanism would still be effective in this case, because players would remain motivated to avoid public bad actions out of concern that their neighbors might sever links. Examining this conjecture, both theoretically and experimentally, is a crucial direction for future research.

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